

# Statistics 210B Lecture 6 Notes

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February 3, 2022

## 1 Gaussian Concentration

### 1.1 Freedman's inequality

Last time, we generalized the Hoeffding and Bernstein inequalities for independent random variables to Azuma-Hoeffding and “Azuma Bernstein inequalities for martingales.”

Our “Azuma-Bernstein” inequality says that if  $\mathbb{E}[e^{\lambda D_k} \mid \mathcal{F}_{k-1}] \leq e^{\lambda^2 \nu_k^2 / 2}$ , then

$$\left| \frac{1}{n} \sum_{k=1}^n D_k \right| \leq \max \left\{ \sqrt{\frac{\frac{2}{n} \sum_{k=1}^n \nu_k^2}{n} \log \left( \frac{2}{\delta} \right)}, \frac{2\alpha_* \log \left( \frac{2}{\delta} \right)}{n} \right\} \quad \text{with probability } 1 - \delta.$$

However, sometimes  $\nu_k^2$  is not deterministic and  $\nu_k^2 = \mathbb{E}[D_k^2 \mid \mathcal{F}_{k-1}]$  instead is  $\mathcal{F}_{k-1}$  measurable.

**Theorem 1.1** (Freedman's inequality). *Let  $\{(D_k, \mathcal{F}_k)\}$  be a martingale difference sequence such that*

1.  $\mathbb{E}[D_k \mid \mathcal{F}_{k-1}] = 0$ .

2.  $D_k \leq b$  a.s.

Then for all  $\lambda \in (0, 1/b)$  and  $\delta \in (0, 1)$ ,

$$\mathbb{P} \left( \sum_{t=1}^T X_t \leq \lambda \sum_{t=1}^T \mathbb{E}[D_k^2 \mid \mathcal{F}_{k-1}] + \frac{\log(1/\delta)}{\lambda} \right) \geq 1 - \delta.$$

This is useful in bandit and reinforcement learning research.<sup>1</sup>

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<sup>1</sup>For example, see Theorem 1 in Beygelzimer, Langford, et. al. 2010.

## 1.2 Maximal Azuma-Hoeffding inequality

Recall Doob's maximal inequality for sub-martingales.

**Lemma 1.1** (Doob's maximal inequality). *If  $\{X_s\}_{s \geq 0}$  is a sub-martingale, i.e.*

$$X_s \leq \mathbb{E}[X_t \mid \mathcal{F}_s] \quad \forall s < t,$$

*then for all  $u > 0$ ,*

$$\mathbb{P} \left( \sup_{0 \leq t \leq T} X_t \geq u \right) \leq \frac{\mathbb{E}[\max\{X_T, 0\}]}{u}.$$

This gives rise to a maximal version of the Azuma-Hoeffding inequality:

**Theorem 1.2** (Maximal Azuma-Hoeffding inequality). *Let  $\{(D_k, \mathcal{F}_k)\}$  be a martingale difference sequence, and suppose there exists  $\{(a_k, b_k)\}_{k=1}^n$  such that  $D_k \in (a_k, b_k)$  a.s. Then*

$$\mathbb{P} \left( \sup_{0 \leq k \leq n} \sum_{s=1}^k D_k \geq t \right) \leq \exp \left( - \frac{2t^2}{\sum_{k=1}^n (b_k - a_k)^2} \right).$$

If we used the usual Azuma-Hoeffding inequality instead, we would need to use a union bound, which would give a factor of  $n$  in the bound. We can write this conclusion as

$$\sup_{0 \leq k \leq n} \sum_{s=1}^k D_k \leq \sqrt{\frac{C \log(1/\delta)}{n}}.$$

If we have the extra factor of  $n$ , we get an  $n/\delta$  instead, which can sometimes be not a big deal for our bound since we are taking a log.

## 1.3 Gaussian concentration

**Lemma 1.2.** *Let  $X_1, X_2, \dots, X_n \stackrel{\text{iid}}{\sim} N(0, 1)$  and  $f : \mathbb{R}^n \rightarrow \mathbb{R}$  such that  $f$  is  $L$ -Lipschitz in  $\|\cdot\|_2$ , i.e.*

$$|f(x) - f(y)| \leq L \|x - y\|_2 \quad \forall x, y \in \mathbb{R}^n.$$

*Then*

1.  $f(X_{1:n}) - \mathbb{E}[f(X_{1:n})]$  is sG( $L$ ).

- 2.

$$\mathbb{P}(|f(X_{1:n}) - \mathbb{E}[f(X_{1:n})]| \geq t) \leq 2 \exp \left( - \frac{t^2}{2L^2} \right).$$

**Remark 1.1.** We need  $f$  to be Lipschitz as a whole function! It's not just sufficient for the function to be coordinate-wise Lipschitz.

**Remark 1.2.** If the  $X_i$ s are non-Gaussian, this doesn't always hold with only Lipschitz-ness.

There are many different proofs of this lemma, but none are very simple.

Proof 1: Gaussian interpolation method

Proof 2: Gaussian isoperimetric inequality

Proof 3: Gaussian log-Sobolev inequality + Herbst argument

Today, we will present a proof using the Gaussian interpolation method, which is useful in research. However, this is a technique where you need to develop some intuition to understand it.

## 1.4 Examples of Gaussian concentration

**Example 1.1** (Order statistics). Let  $(X_i)_{i \in [n]} \stackrel{\text{iid}}{\sim} N(0, 1)$ . The order statistics are the random variables arranged in increasing order:  $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$ . Let  $f_k(X_{1:n}) = X_{(k)}$ . This is Lipschitz:

$$\begin{aligned} |f_k(X_{1:n}) - f_k(Y_{1:n})| &= |X_{(k)} - Y_{(k)}| \\ &\leq \sqrt{\sum_{k=1}^n |X_{(k)} - Y_{(k)}|^2} \end{aligned}$$

The **rearrangement inequality** says that if you sort the terms, the distance is greater than the distance of with unsorted terms.

$$\begin{aligned} &\leq \sqrt{\sum_{k=1}^n |X_k - Y_k|^2} \\ &= \|X - Y\|_2. \end{aligned}$$

This means that  $L = 1$ , so  $X_{(k)} - \mathbb{E}[X_{(k)}]$  is  $\text{sG}(1)$ . Therefore,

$$|X_{(k)} - \mathbb{E}[X_{(k)}]| \leq \sqrt{\log(2/\delta)} \quad \text{with probability } 1 - \delta.$$

If we apply this to  $k = n$ , we get

$$\left| \max_{i \in [n]} X_i - \underbrace{\mathbb{E} \left[ \max_{i \in [n]} X_i \right]}_{\sqrt{2 \log n}} \right| = O_p(1).$$

**Example 1.2** (Singular value of Gaussian random matrices). Let

$$X = \begin{bmatrix} X_{1,1} & \cdots & X_{1,d} \\ \vdots & & \vdots \\ X_{n,1} & \cdots & X_{n,d} \end{bmatrix} \in \mathbb{R}^{n \times d}, \quad X_{i,j} \stackrel{\text{iid}}{\sim} N(0, 1).$$

Let  $f_k(X) = \sigma_k(X)$  be the  $k$ -th largest singular value of  $X$ . For example,  $f_1(X) = \|X\|_{\text{op}}$ . It can be shown that  $\mathbb{E}[\|X\|_{\text{op}}] \approx \sqrt{n} + \sqrt{d}$ . We can show that  $f_k$  is Lipschitz; what is the norm we want to be using for a matrix? Define the vectorized version of the matrix as  $\text{vec}(X) := (X_{1,1}, X_{1,2}, \dots, X_{1,d}, X_{2,1}, \dots, X_{2,d}, \dots, X_{n,d})$ . Then

$$\|\text{vec}(X) - \text{vec}(Y)\|_2 = \|X - Y\|_F = \sqrt{\sum_{i,j} (X_{i,j} - Y_{i,j})^2},$$

where  $\|\cdot\|_F$  is the **Frobenius norm**. Now we have

$$|f_k(X) - f_k(Y)| \leq |\sigma_k(X) - \sigma_k(Y)|$$

**Weyl's inequality**, a deterministic linear algebra result, says that

$$\begin{aligned} &\leq \|X - Y\|_{\text{op}} \\ &\leq \|X - Y\|_F, \end{aligned}$$

so  $L = 1$ . Weyl's inequality can be proven by using the variational representation of singular values.

This calculation tells us that  $f_k(X) - \mathbb{E}[f_k(X)]$  is sG(1), so

$$f_k(X) - \mathbb{E}[f_k(X)] \leq \sqrt{\log(2/\delta)} \quad \text{with probability } 1 - \delta.$$

Applying this to  $k = 1$  gives

$$|\|X\|_{\text{op}} - \underbrace{\mathbb{E}[\|X\|_{\text{op}}]}_{\sqrt{n} + \sqrt{d}}| = O(1).$$

## 1.5 Gaussian complexity

Gaussian complexity is a very important notion in compressed sensing. Suppose we have a set  $A \subseteq \mathbb{R}^n$ . How do we measure its “size”? A reasonable size function  $S$  should at least satisfy  $S(A) \leq S(B)$  if  $A \subseteq B$ . Here are some reasonable size functions:

1. Euclidean width:  $D(A) = \max_{a \in A} \|a\|_2$ .
2. Dimension: A line has dimension 1, and a plane has dimension 2.

**Definition 1.1.** Given a set  $A$ , let  $W = (W_1, \dots, W_n)^\top \in \mathbb{R}^n$  with  $W_i \stackrel{\text{iid}}{\sim} N(0, 1)$ . The **Gaussian complexity** or "statistical dimension" of  $A$  is

$$\mathcal{G}(A) := \mathbb{E}_{W \sim N(0, I_n)} \left[ \sup_{a \in A} \langle a, W \rangle \right].$$

Note that if we don't take the supremum in the expectation, the quantity would be 0. This quantity is always nonnegative.

**Example 1.3.** Let  $B_p(r) = \{x \in \mathbb{R}^n : \|x\|_p \leq r\}$ . Then

$$\mathcal{G}(B_p(r)) = \mathbb{E} \left[ \sup_{\|x\|_p \leq r} \langle x, W \rangle \right]$$

If  $q$  is the conjugate exponent of  $p$ , so  $\frac{1}{p} + \frac{1}{q} = 1$ , this is the variational representation of the  $\|\cdot\|_q$  norm:

$$\begin{aligned} r \mathbb{E}[\|W\|_q] \\ \approx rn^{1/q}. \end{aligned}$$

Note that if  $p_1 \leq p_2$ , then  $q_1 \geq q_2$ , so  $\mathcal{G}(B_{p_1}(r)) \leq \mathcal{G}(B_{p_2}(r))$ .

We want to show that  $f(W) := \sup_{a \in A} \langle a, W \rangle$  concentrates. Fix  $w, w' \in \mathbb{R}^n$ . Then

$$f(w) - f(w') = \sup_{a \in A} \langle a, w \rangle - \sup_{a \in A} \langle a, w' \rangle$$

Denote  $a^* = \arg \max_a \langle a, w \rangle$

$$\begin{aligned} &= \langle a^*, w \rangle - \sup_{a \in A} \langle a, w' \rangle \\ &= \inf_{a \in A} \langle a^*, w \rangle - \langle a, w' \rangle \\ &\leq \langle a^* w - w' \rangle \\ &\leq \|a^*\| \|w - w'\|_2 \\ &\leq D(A) \|w - w'\|_2. \end{aligned}$$

The other side can be proven similarly, so  $f$  is  $D(A)$ -Lipschitz. Concentration says that  $f(W)$  is  $\text{sG}(D(A))$ .

**Example 1.4.** If we let  $A = B_2(R)$ , then

$$\mathbb{E}[f(W)] = \mathcal{G}(B_2(r)) = r\sqrt{n},$$

since  $D(A) = r$ .

## 1.6 Proof of the Gaussian concentration inequality (interpolation method)

**Lemma 1.3.** *For all convex  $\phi : \mathbb{R} \rightarrow \mathbb{R}$  and differentiable  $f : \mathbb{R}^n \rightarrow \mathbb{R}$ ,*

$$\mathbb{E}[\phi(f(X) - \mathbb{E}[f(Y)])] \leq \mathbb{E}[\phi(\frac{\pi}{2} \langle \nabla f(X), Y \rangle)],$$

where  $X, Y \stackrel{\text{iid}}{\sim} N(0, I_n)$ .

First, assume this lemma holds, and prove Gaussian concentration:

*Proof.* Take  $\phi = \exp(\lambda \cdot)$ . Then

$$\mathbb{E}[\exp(\lambda(f(X) - \mathbb{E}[f(Y)]))] \leq \mathbb{E}[\exp(\lambda \frac{\pi}{2} \langle \nabla f(X), Y \rangle)]$$

Observe that  $\frac{\pi}{2} \langle \nabla f(X), Y \rangle$  is  $N(0, \frac{\pi^2}{4} \|\nabla f(X)\|_2^2)$  given  $X$ .

$$\begin{aligned} &= \mathbb{E}_X[\exp(\frac{\lambda^2}{2} \frac{\pi^2}{4} \|\nabla f(X)\|_2^2)] \\ &\leq \exp\left(\frac{\lambda^2}{2} \frac{\pi^2}{4} L^2\right). \end{aligned}$$

This says that  $f(X) - \mathbb{E}[f(X)]$  is  $\text{sG}(\frac{\pi}{2}L)$ . □

The above proof gives a worse constant, but the constant can be improved with different methods. Here is the proof of the lemma:

*Proof.* First, use conditioning and Jensen's inequality to say that.

$$\mathbb{E}[\phi(f(X) - \mathbb{E}[f(Y)])] \leq \mathbb{E}_{X,Y}[\phi(f(X) - f(Y))]$$

The idea is to use the integral representation of the Taylor expansion to interpolate between  $X$  and  $Y$ . Observe that if  $Z(\theta) = X \cos \theta + Y \sin \theta$ , then for every  $\theta$ ,  $Z(\theta) \stackrel{d}{=} X \stackrel{d}{=} Y$  and  $Z'(\theta) \stackrel{d}{=} X \stackrel{d}{=} Y$ . Another important property is that  $Z(\theta) \perp Z'(\theta)$ ; this is because  $Z(\theta), Z'(\theta)$  are Gaussians with 0 covariance. Now

$$f(X) - f(Y) = \int_0^{\pi/2} \langle \nabla f(Z(\theta)), Z'(\theta) \rangle d\theta,$$

so we can write

$$\mathbb{E}[\phi(f(X) - f(Y))] = \mathbb{E}\left[\phi\left(\int_0^{\pi/2} \langle \nabla f(Z(\theta)), Z'(\theta) \rangle d\theta\right)\right]$$

Using Jensen's inequality,  $\phi(\int \cdot d\theta) \leq \int \phi(\cdot) d\theta$  when  $\int \cdot d\theta = 1$ .

$$\begin{aligned} &\leq \frac{2}{\pi} \int_0^{\pi/2} \mathbb{E}[\phi(\frac{\pi}{2} \langle \nabla f(Z(\theta)), Z'(\theta) \rangle)] d\theta \\ &= \mathbb{E}[\phi(\frac{\pi}{2} \langle \nabla f(X), Y \rangle)]. \end{aligned} \quad \square$$

This proof is very delicate, and the construction looks ad hoc, but it is actually very useful in a variety of situations.

## 1.7 Other methods for establishing concentration

1. Matrix concentration: If  $(X_i)_{i \in [n]} \subseteq \mathbb{R}^{m \times d}$  with  $X_i \stackrel{\text{iid}}{\sim} X$ , can we find a bound for

$$\left\| \frac{1}{n} \sum_{i=1}^n X_i - \mathbb{E}[X_i] \right\|_{\text{op}} ?$$

The answer is yes; there is a matrix Bernstein inequality, Rudelson's inequality, and a matrix Freedman inequality. These involve the matrix MGF and Lieb's inequality. For more, see *An Introduction to Matrix Inequalities*, Tropp 2015, and *Introduction to Non-asymptotic analysis of random matrices*, Vershynin 2010.

2. Entropy method and the Herbst argument

**Definition 1.2.** The **Herbst** argument is that a sufficient condition for  $X$  to be  $\text{sG}(\sigma)$  is to show that

$$\mathbb{H}(e^{\lambda X}) \leq \frac{\lambda^2 \sigma^2}{2} \mathbb{E}[e^{\lambda X}],$$

where  $\mathbb{H}$  is the entropy.

Why do we want to look at  $\mathbb{H}(e^{\lambda X})$ ? This is because it has a good **tensorization property** when  $X_i$  are independent:

$$\mathbb{H}(e^{\lambda f(X_{1:n})}) \leq \mathbb{E} \left[ \sum_{i=1}^n \mathbb{H} \left( \underbrace{e^{\lambda f_k(X_k)}}_{\text{easy to handle when } f_k \text{ Lip., } X_k \text{ bdd.}} \mid X^{\setminus k} \right) \right]$$

For this, see chapter 3.1 of Wainwright's textbook or chapter 3 of van Handel's textbook

3. Isoperimetric inequality: This is a geometric property in  $\mathbb{R}^n$  with Lebesgue measure. If  $A \subseteq \mathbb{R}^n$  has fixed volume and we want to minimize the perimeter, then the solution is when  $A$  is a ball. This generalizes to other measures:

$X \sim \mu =$	$N(0, I_n)$	$S^{n-1}(\sqrt{n})$	$\text{Unif}(\{\pm 1\}^n)$
	Half space	Spherical cap	Hamming ball

The isoperimetric inequality implies that  $f(X)$  concentrates when  $f$  is Lipschitz. For this, see chapter 3.2 of Wainwright's book and also see Chapter 7 of the book by Lugosi, Massart, and Boucheron.

4. Transportation approach:

**Lemma 1.4** (Bobkov-Gotze). *Given a measure  $\mu \in \mathcal{P}(\mathbb{R}^n)$ ,*

$$X \sim \mu, \forall f \text{ 1-Lipschitz, } f(X) \text{ is sG}(\sigma) \iff W_1(\nu, \mu) \leq \sqrt{2\sigma^2 \text{KL}(\nu \parallel \mu)} \forall \nu \in \mathcal{P}(\mathbb{R}^n),$$

*where  $W_1$  is the transportation distance and  $\text{KL}$  is the relative entropy.*

This property on the right also tensorizes in some way. For more on this, see chapter 3.3 in Wainwright's book or chapter 4 in van Handel's book.